

Week 5: Simple Linear Regression

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¹These slides are heavily influenced by Matt Blackwell, Adam Glynn and Jens Hainmueller. Illustrations by Shay O'Brien.

Where We've Been and Where We're Going...

- Last Week
 - ▶ hypothesis testing
 - ▶ what is regression
- This Week
 - ▶ Monday:
 - ★ mechanics of OLS
 - ★ properties of OLS
 - ▶ Wednesday:
 - ★ hypothesis tests for regression
 - ★ confidence intervals for regression
 - ★ goodness of fit
- Next Week
 - ▶ mechanics with two regressors
 - ▶ omitted variables, multicollinearity
- Long Run
 - ▶ probability $\rightarrow$ inference $\rightarrow$ regression $\rightarrow$ causal inference

Questions?

Macrostructure

The next few weeks,

- Linear Regression with Two Regressors
- Multiple Linear Regression
- Break Week
- What Can Go Wrong and How to Fix It
- Regression in the Social Sciences and Introduction to Causality
- Thanksgiving
- Causality with Measured Confounding
- Unmeasured Confounding and Instrumental Variables
- Repeated Observations and Panel Data

Review session timing.

- 1 Mechanics of OLS
- 2 Properties of the OLS estimator
- 3 Example and Review
- 4 Properties Continued
- 5 Hypothesis tests for regression
- 6 Confidence intervals for regression
- 7 Goodness of fit
- 8 Wrap Up of Univariate Regression
- 9 Fun with Non-Linearities
- 10 Appendix: r^2 derivation

The population linear regression function

- The (population) simple linear regression model can be stated as the following:

$$r(x) = E[Y|X = x] = \beta_0 + \beta_1 x$$

- This (partially) describes the **data generating process** in the population
- Y = dependent variable
- X = independent variable
- β_0, β_1 = population intercept and population slope (what we want to estimate)

The sample linear regression function

- The **estimated** or sample regression function is:

$$\widehat{r}(X_i) = \widehat{Y}_i = \widehat{\beta}_0 + \widehat{\beta}_1 X_i$$

- $\widehat{\beta}_0, \widehat{\beta}_1$ are the estimated intercept and slope
- $\widehat{Y}_i$ is the fitted/predicted value
- We also have the residuals, $\widehat{u}_i$ which are the differences between the true values of Y and the predicted value:

$$\widehat{u}_i = Y_i - \widehat{Y}_i$$

- You can think of the residuals as the prediction errors of our estimates.

Overall Goals for the Week

- Learn how to run and read regression
- **Mechanics**: how to estimate the intercept and slope?
- **Properties**: when are these good estimates?
- **Uncertainty**: how will the OLS estimator behave in repeated samples?
- **Testing**: can we assess the plausibility of no relationship ($\beta_1 = 0$)?
- **Interpretation**: how do we interpret our estimates?

What is OLS?

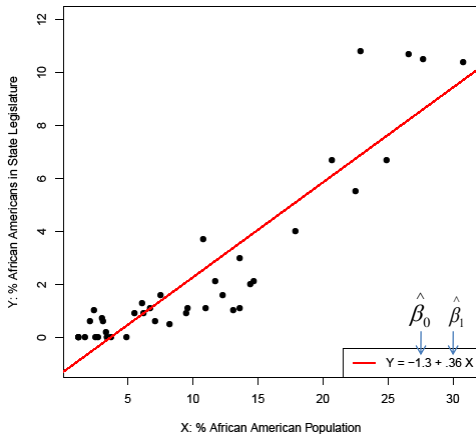
- An estimator for the slope and the intercept of the regression line
- We talked last week about ways to derive this estimator and we settled on deriving it by **minimizing the squared prediction errors** of the regression, or in other words, minimizing the sum of the squared residuals:
- **Ordinary Least Squares (OLS):**

$$(\hat{\beta}_0, \hat{\beta}_1) = \arg \min_{b_0, b_1} \sum_{i=1}^n (Y_i - b_0 - b_1 X_i)^2$$

- In words, the OLS estimates are the intercept and slope that minimize the **sum of the squared residuals**.

Graphical Example

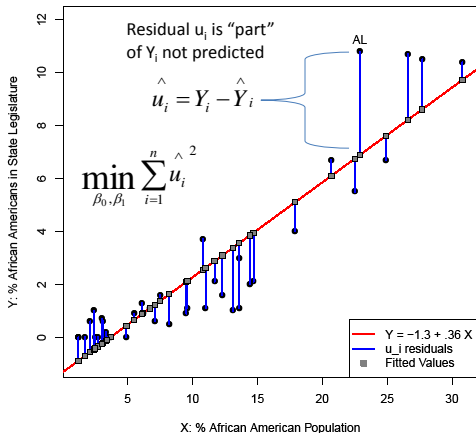
How do we fit the regression line $\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X$ to the data?



Graphical Example

How do we fit the regression line $\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X$ to the data?

Answer: We will **minimize the squared sum of residuals**



Deriving the OLS estimator

- Let's think about n pairs of sample observations:
 $(Y_1, X_1), (Y_2, X_2), \dots, (Y_n, X_n)$
- Let $\{b_0, b_1\}$ be possible values for $\{\beta_0, \beta_1\}$
- Define the **least squares objective function**:

$$S(b_0, b_1) = \sum_{i=1}^n (Y_i - b_0 - b_1 X_i)^2.$$

- How do we derive the LS estimators for β_0 and β_1 ? We want to minimize this function, which is actually a very well-defined calculus problem.
 - 1 Take partial derivatives of S with respect to b_0 and b_1 .
 - 2 Set each of the partial derivatives to 0
 - 3 Solve for $\{b_0, b_1\}$ and replace them with the solutions
- To the board we go!

The OLS estimator

- Now we're done! Here are the **OLS estimators**:

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

Intuition of the OLS estimator

- The intercept equation tells us that the regression line goes through the point $(\bar{Y}, \bar{X})$:

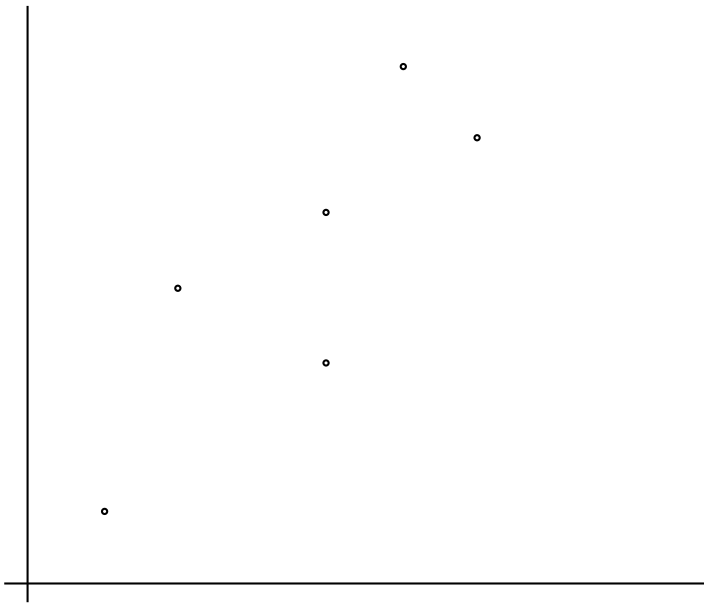
$$\bar{Y} = \hat{\beta}_0 + \hat{\beta}_1 \bar{X}$$

- The slope for the regression line can be written as the following:

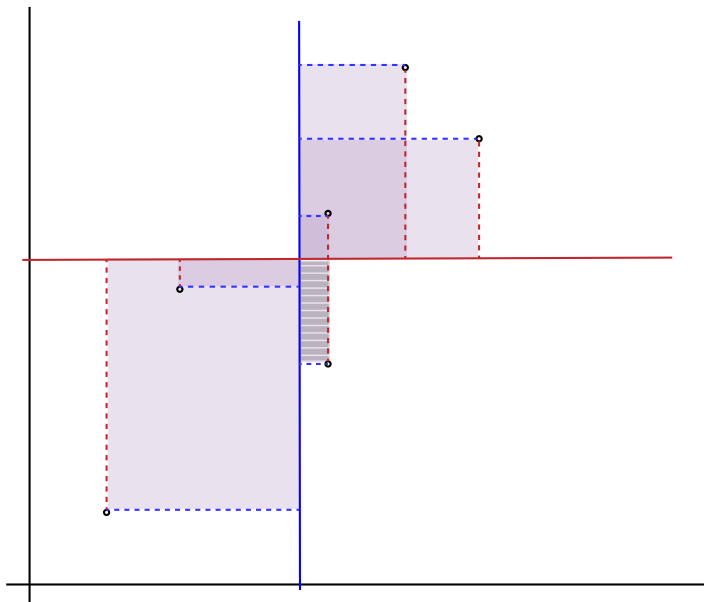
$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\text{Sample Covariance between } X \text{ and } Y}{\text{Sample Variance of } X}$$

- The higher the **covariance** between X and Y , the higher the **slope** will be.
- Negative covariances $\rightarrow$ negative slopes;
positive covariances $\rightarrow$ positive slopes
- What happens when X_i doesn't vary?
- What happens when Y_i doesn't vary?

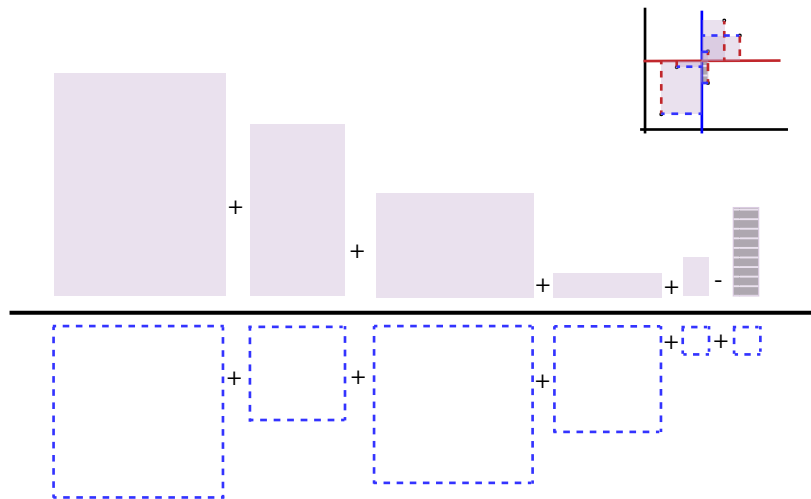
A Visual Intuition for the OLS Estimator



A Visual Intuition for the OLS Estimator



A Visual Intuition for the OLS Estimator



Mechanical properties of OLS

- Later we'll see that under certain assumptions, OLS will have nice statistical properties.
- But some properties are mechanical since they can be derived from the first order conditions of OLS.

- 1 The residuals will be 0 on average:

$$\frac{1}{n} \sum_{i=1}^n \hat{u}_i = 0$$

- 2 The residuals will be uncorrelated with the predictor ($\widehat{\text{cov}}$ is the sample covariance):

$$\widehat{\text{cov}}(X_i, \hat{u}_i) = 0$$

- 3 The residuals will be uncorrelated with the fitted values:

$$\widehat{\text{cov}}(\hat{Y}_i, \hat{u}_i) = 0$$

OLS slope as a weighted sum of the outcomes

- One useful derivation is to write the OLS estimator for the slope as a weighted sum of the outcomes.

$$\hat{\beta}_1 = \sum_{i=1}^n W_i Y_i$$

- Where here we have the weights, W_i as:

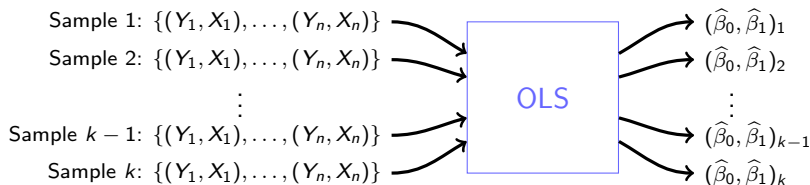
$$W_i = \frac{(X_i - \bar{X})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

- This is important for two reasons. First, it'll make derivations later much easier. And second, it shows that is just the sum of a random variable. Therefore it is also a random variable.
- To the board!

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Sampling distribution of the OLS estimator

- Remember: OLS is an estimator—it's a machine that we plug data into and we get out estimates.

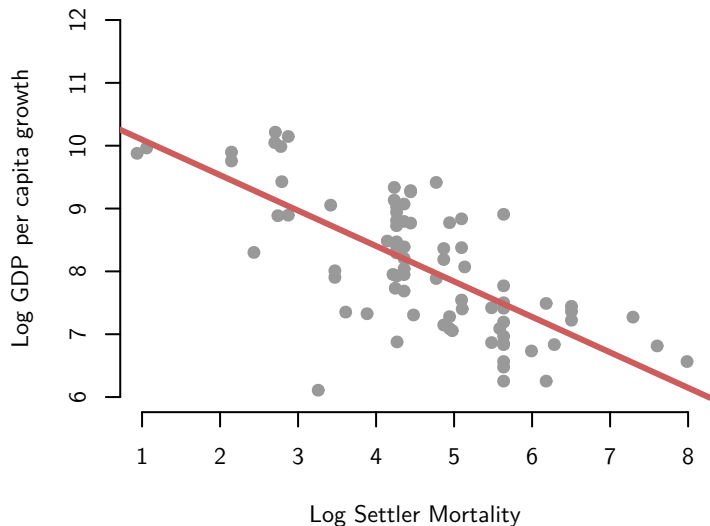


- Just like the sample mean, sample difference in means, or the sample variance
- It has a sampling distribution, with a sampling variance/standard error, etc.
- Let's take a simulation approach to demonstrate:
 - Pretend that the AJR data represents the population of interest
 - See how the line varies from sample to sample

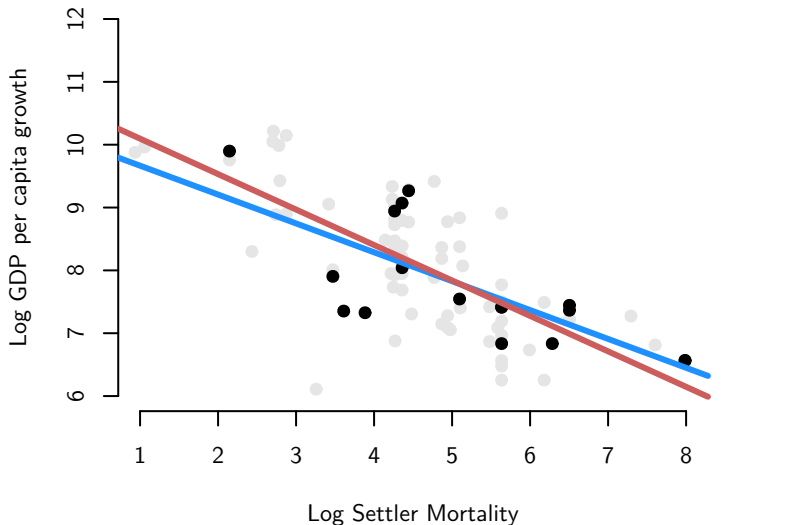
Simulation procedure

- 1 Draw a random sample of size $n = 30$ with replacement using `sample()`
- 2 Use `lm()` to calculate the OLS estimates of the slope and intercept
- 3 Plot the estimated regression line

Population Regression



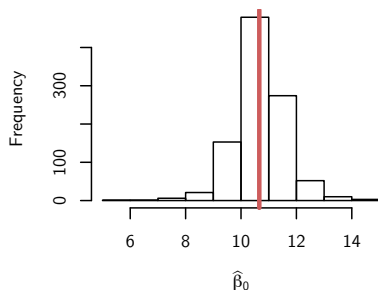
Randomly sample from AJR



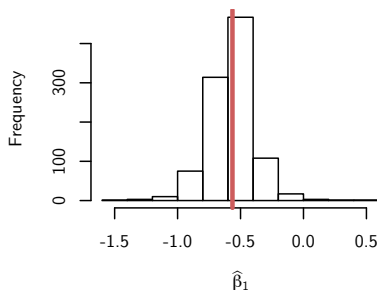
Sampling distribution of OLS

- You can see that the estimated slopes and intercepts vary from sample to sample, but that the “average” of the lines looks about right.

Sampling distribution of intercepts



Sampling distribution of slopes



- Is this unique?

Assumptions for unbiasedness of the sample mean

- What assumptions did we make to prove that the sample mean was unbiased?

$$\mathbb{E}[\bar{X}] = \mu$$

- Just one: random sample
- We'll need more than this for the regression case

Our goal

- What is the sampling distribution of the OLS slope?

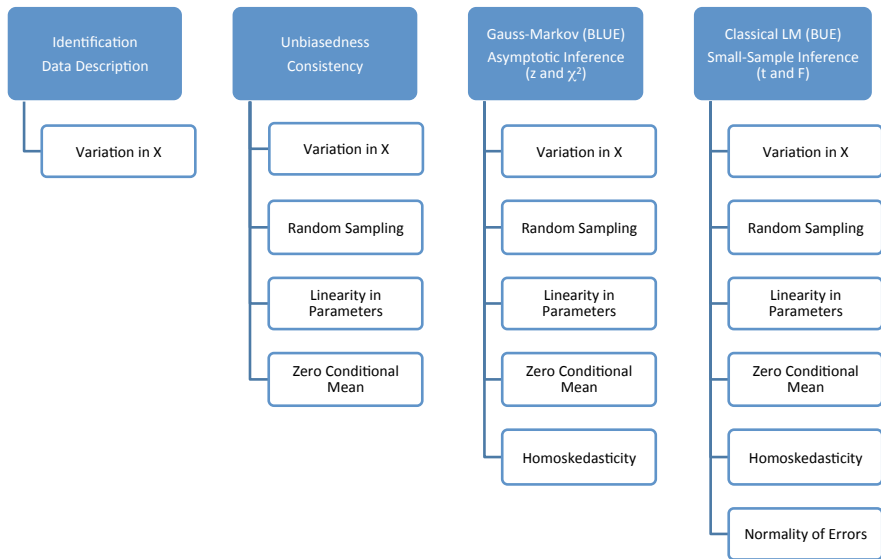
$$\hat{\beta}_1 \sim ?(?, ?)$$

- We need fill in those ?s.
- We'll start with the mean of the sampling distribution. Is the estimator centered at the true value, β_1 ?
- Most of our derivations will be in terms of the slope but they apply to the intercept as well.

OLS Assumptions Preview

- 1 **Linearity in Parameters:** The population model is linear in its parameters and correctly specified
- 2 **Random Sampling:** The observed data represent a random sample from the population described by the model.
- 3 **Variation in X :** There is variation in the explanatory variable.
- 4 **Zero conditional mean:** Expected value of the error term is zero conditional on all values of the explanatory variable
- 5 **Homoskedasticity:** The error term has the same variance conditional on all values of the explanatory variable.
- 6 **Normality:** The error term is independent of the explanatory variables and normally distributed.

Hierarchy of OLS Assumptions



OLS Assumption I

Assumption (I. Linearity in Parameters)

The population regression model is linear in its parameters and correctly specified as:

$$Y = \beta_0 + \beta_1 X_1 + u$$

- Note that it can be nonlinear *in variables*
 - ▶ OK: $Y = \beta_0 + \beta_1 X + u$ or
 $Y = \beta_0 + \beta_1 X^2 + u$ or
 $Y = \beta_0 + \beta_1 \log(X) + u$
 - ▶ Not OK: $Y = \beta_0 + \beta_1^2 X + u$ or
 $Y = \beta_0 + \exp(\beta_1) X + u$
- β_0, β_1 : Population **parameters** — fixed and unknown
- u : Unobserved random variable with $E[u] = 0$ — captures all other factors influencing Y other than X
- We assume this to be the structural model, i.e., the model describing the true process generating Y

OLS Assumption II

Assumption (II. Random Sampling)

The observed data:

$$(y_i, x_i) \text{ for } i = 1, \dots, n$$

represent an i.i.d. random sample of size n following the population model.

Data examples consistent with this assumption:

- A cross-sectional survey where the units are sampled randomly

Potential Violations:

- Time series data (regressor values may exhibit persistence)
- Sample selection problems (sample not representative of the population)

OLS Assumption III

Assumption (III. Variation in X ; a.k.a. No Perfect Collinearity)

The observed data:

$$x_i \text{ for } i = 1, \dots, n$$

are not all the same value.

Satisfied as long as there is some variation in the regressor X in the sample.

Why do we need this?

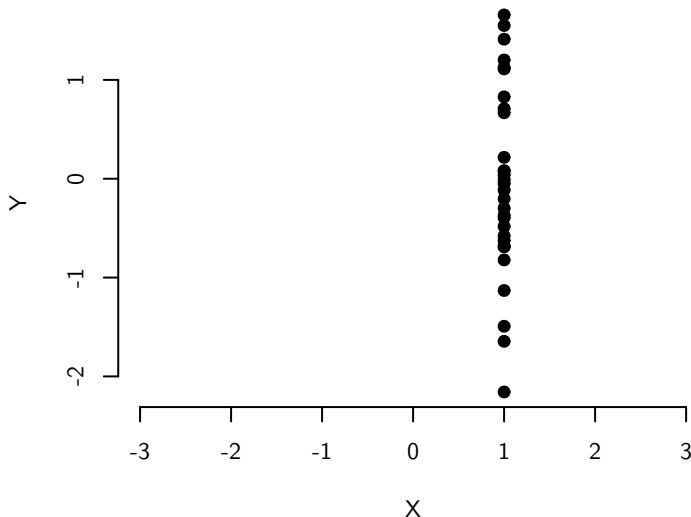
$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

This assumption is needed just to calculate $\hat{\beta}$.

Only assumption needed for using OLS as a pure data summary.

Stuck in a moment

- Why does this matter? How would you draw the line of best fit through this scatterplot, which is a violation of this assumption?



OLS Assumption IV

Assumption (IV. Zero Conditional Mean)

The expected value of the error term is zero conditional on any value of the explanatory variable:

$$E[u|X] = 0$$

- $E[u|X] = 0$ implies a slightly weaker condition $\text{Cov}(X, u) = 0$
- Given random sampling, $E[u|X] = 0$ also implies $E[u_i|x_i] = 0$ for all i

How does this assumption get violated? Let's generate data from the following model:

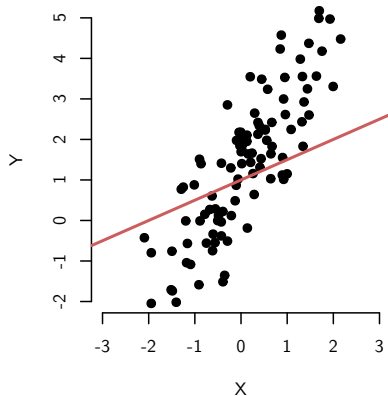
$$Y_i = 1 + 0.5X_i + u_i$$

But let's compare two situations:

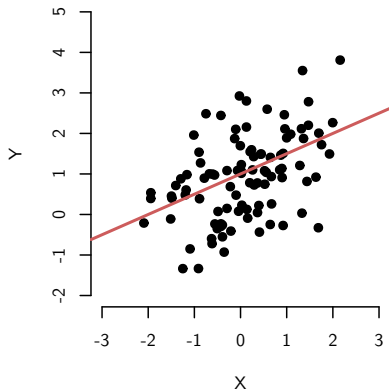
- 1 Where the mean of u_i depends on X_i (they are correlated)
- 2 No relationship between them (satisfies the assumption)

Violating the zero conditional mean assumption

Assumption 4 violated

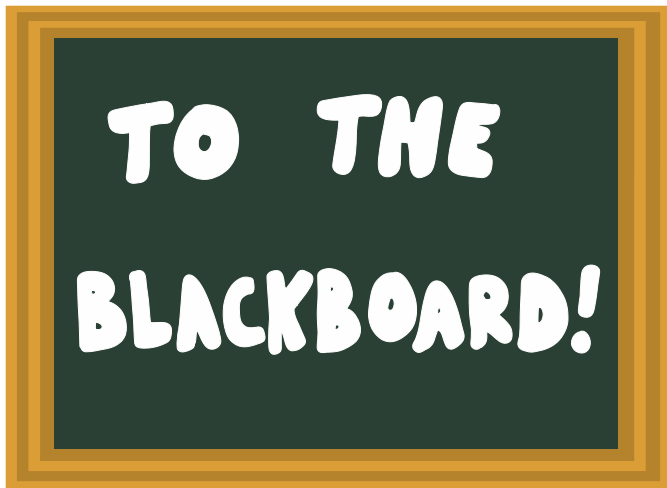


Assumption 4 not violated



Unbiasedness (to the blackboard)

With Assumptions 1-4, we can show that the OLS estimator for the slope is unbiased, that is $E[\hat{\beta}_1] = \beta_1$.



Unbiasedness of OLS

Theorem (Unbiasedness of OLS)

Given OLS Assumptions I–IV:

$$E[\hat{\beta}_0] = \beta_0 \quad \text{and} \quad E[\hat{\beta}_1] = \beta_1$$

The sampling distributions of the estimators $\hat{\beta}_1$ and $\hat{\beta}_0$ are centered about the true population parameter values β_1 and β_0 .

Where are we?

- Now we know that, under Assumptions 1-4, we know that

$$\hat{\beta}_1 \sim ?(\beta_1, ?)$$

- That is we know that the sampling distribution is **centered on the true population slope**, but we don't know the population variance.

Sampling variance of estimated slope

- In order to derive the sampling variance of the OLS estimator,
 - 1 Linearity
 - 2 Random (iid) sample
 - 3 Variation in X_i
 - 4 Zero conditional mean of the errors
 - 5 Homoskedasticity

Variance of OLS Estimators

How can we derive $\text{Var}[\hat{\beta}_0]$ and $\text{Var}[\hat{\beta}_1]$? Let's make the following additional assumption:

Assumption (V. Homoskedasticity)

The conditional variance of the error term is constant and does not vary as a function of the explanatory variable:

$$\text{Var}[u|X] = \sigma_u^2$$

- This implies $\text{Var}[u] = \sigma_u^2$
→ all errors have an identical **error variance** ($\sigma_{u_i}^2 = \sigma_u^2$ for all i)
- Taken together, Assumptions I–V imply:

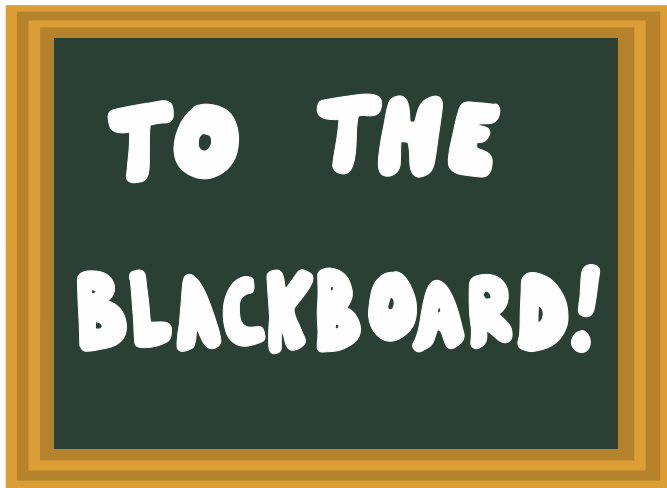
$$E[Y|X] = \beta_0 + \beta_1 X$$

$$\text{Var}[Y|X] = \sigma_u^2$$

- Violation: $\text{Var}[u|X = x_1] \neq \text{Var}[u|X = x_2]$ called **heteroskedasticity**.
- Assumptions I–V are collectively known as the **Gauss-Markov assumptions**

Deriving the sampling variance

$$\text{var}[\hat{\beta}_1 | X_1, \dots, X_n] = ??$$



Variance of OLS Estimators

Theorem (Variance of OLS Estimators)

Given OLS Assumptions I–V (Gauss-Markov Assumptions):

$$\text{Var}[\hat{\beta}_1 | X] = \frac{\sigma_u^2}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

$$\text{Var}[\hat{\beta}_0 | X] = \sigma_u^2 \left\{ \frac{1}{n} + \frac{\bar{x}^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \right\}$$

where $\text{Var}[u | X] = \sigma_u^2$ (the error variance).

Understanding the sampling variance

$$\text{var}[\hat{\beta}_1 | X_1, \dots, X_n] = \frac{\sigma_u^2}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

- What drives the sampling variability of the OLS estimator?
 - ▶ The higher the variance of $Y_i | X_i$, the higher the sampling variance
 - ▶ The lower the variance of X_i , the higher the sampling variance
 - ▶ As we increase n , the denominator gets large, while the numerator is fixed and so the sampling variance shrinks to 0.
- But, this formula depends upon an unobserved term: σ_u^2

Estimating the Variance of OLS Estimators

How can we estimate the unobserved error variance $\text{Var}[u] = \sigma_u^2$?

We can derive an estimator based on the **residuals**:

$$\hat{u}_i = y_i - \hat{y}_i = y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i$$

Recall: The **errors** u_i are NOT the same as the residuals $\hat{u}_i$.

Intuitively, the scatter of the residuals around the fitted regression line should reflect the unseen scatter about the true population regression line.

We can measure scatter with the mean squared deviation:

$$MSD(\hat{u}) \equiv \frac{1}{n} \sum_{i=1}^n (\hat{u}_i - \bar{\hat{u}})^2 = \frac{1}{n} \sum_{i=1}^n \hat{u}_i^2$$

Intuitively, which line is likely to be closer to the observed sample values on X and Y , the true line $y_i = \beta_0 + \beta_1 x_i$ or the fitted regression line $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$?

Estimating the Variance of OLS Estimators

- By construction, the regression line is closer since it is drawn to fit the sample we observe
- Specifically, the regression line is drawn so as to minimize the sum of the squares of the distances between it and the observations
- So the spread of the residuals $MSD(\hat{u})$ will slightly *underestimate* the error variance $\text{Var}[u] = \sigma_u^2$ on average
- In fact, we can show that with a single regressor X we have:

$$E[MSD(\hat{u})] = \frac{n-2}{n} \sigma_u^2 \text{ (degrees of freedom adjustment)}$$

- Thus, an **unbiased estimator** for the error variance is:

$$\hat{\sigma}_u^2 = \frac{n}{n-2} MSD(\hat{u}) = \frac{n}{n-2} \frac{1}{n} \sum_{i=1}^n \hat{u}_i^2 = \frac{1}{n-2} \sum_{i=1}^n \hat{u}_i^2$$

We plug this estimate into the variance estimators for $\hat{\beta}_0$ and $\hat{\beta}_1$.

Where are we?

- Under Assumptions 1-5, we know that

$$\hat{\beta}_1 \sim? \left(\beta_1, \frac{\sigma_u^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right)$$

- Now we know the mean and sampling variance of the sampling distribution.
- Next Time: how does this compare to other estimators for the population slope?

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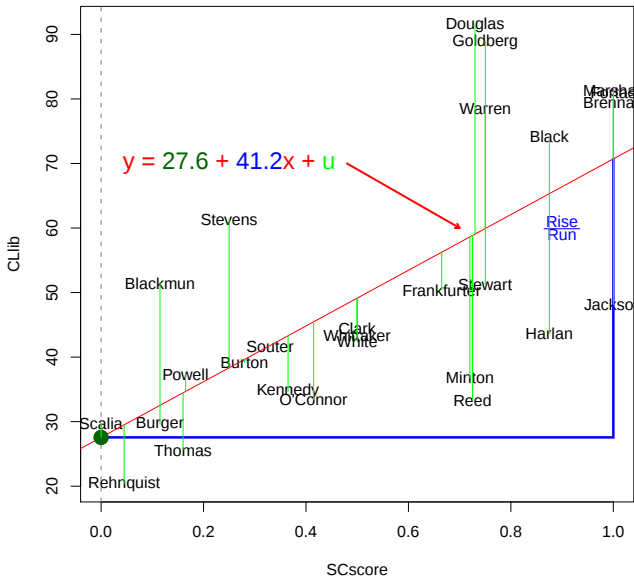
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Questions?

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Example: Epstein and Mershon SCOTUS data

- Data on 27 justices from the Warren, Burger, and Rehnquist courts (can be interpreted as a **census**)
- Percentage of votes in liberal direction for each justice in a number of issue areas
- Segal-Cover scores for each justice
- Party of appointing president



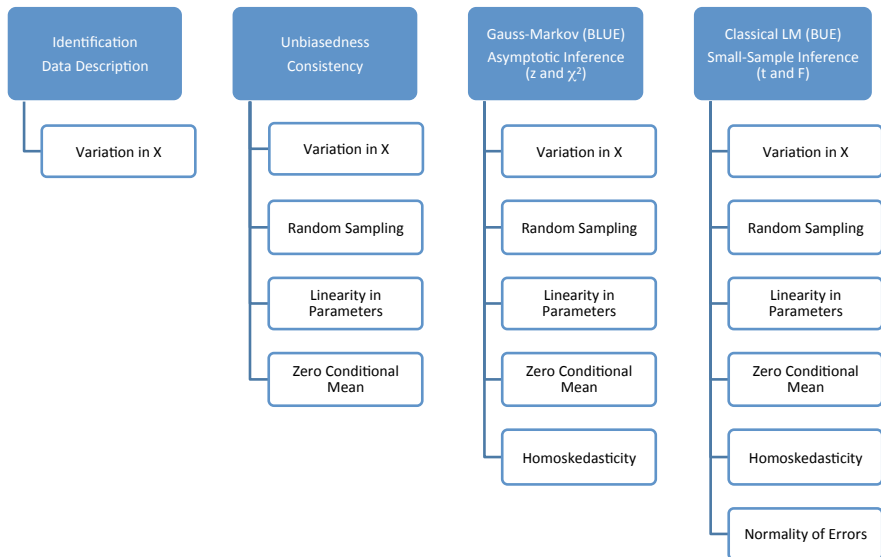
How to get β_0 and β_1

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}.$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}.$$

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Where are we?



Where are we?

- Under Assumptions 1-5, we know that

$$\hat{\beta}_1 \sim? \left(\beta_1, \frac{\sigma_u^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right)$$

- Now we know the mean and sampling variance of the sampling distribution.
- How does this compare to other estimators for the population slope?

OLS is BLUE :(

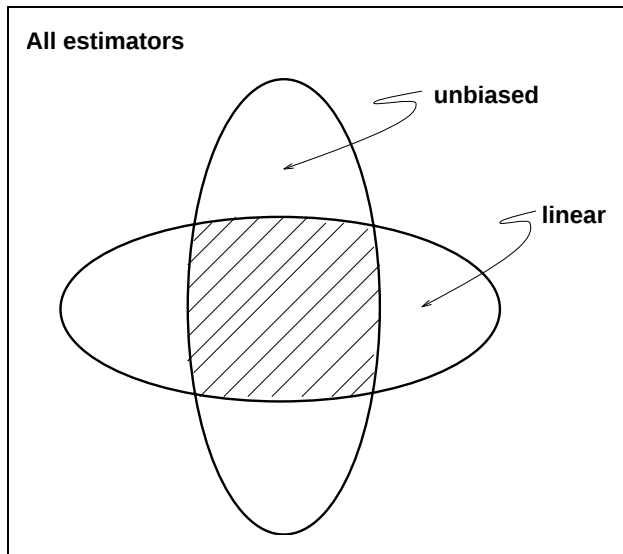
Theorem (Gauss-Markov)

Given OLS Assumptions I–V, the OLS estimator is **BLUE**, i.e. the

- 1 **B**est: Lowest variance in class
- 2 **L**inear: Among Linear estimators
- 3 **U**nbiased: Among Linear Unbiased estimators
- 4 **E**stimator.

- Assumptions 1-5: the “Gauss Markov Assumptions”
- The proof is detailed and doesn't yield insight, so we skip it. (We will explore the intuition some more in a few slides)
- Fails to hold when the assumptions are violated!

Gauss-Markov Theorem



Where are we?

- Under Assumptions 1-5, we know that

$$\hat{\beta}_1 \sim ? \left(\beta_1, \frac{\sigma_u^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right)$$

- And we know that $\frac{\sigma_u^2}{\sum_{i=1}^n (X_i - \bar{X})^2}$ is the lowest variance of any linear estimator of β_1
- What about the last question mark? What's the form of the distribution? Uniform? t ? Normal? Exponential? Hypergeometric?

Large-sample distribution of OLS estimators

- Remember that the OLS estimator is the sum of independent r.v.'s:

$$\hat{\beta}_1 = \sum_{i=1}^n W_i Y_i$$

- Mantra of the Central Limit Theorem:

"the sums and means of r.v.'s tend to be Normally distributed in large samples."

- True here as well, so we know that in large samples:

$$\frac{\hat{\beta}_1 - \beta_1}{SE[\hat{\beta}_1]} \sim N(0, 1)$$

- Can also replace SE with an estimate:

$$\frac{\hat{\beta}_1 - \beta_1}{\widehat{SE}[\hat{\beta}_1]} \sim N(0, 1)$$

Where are we?

Under Assumptions 1-5 and in large samples, we know that

$$\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma_u^2}{\sum_{i=1}^n (X_i - \bar{X})^2}\right)$$



Sampling distribution in small samples

- What if we have a small sample? What can we do then?
- Can't get something for nothing, but we can make progress if we make another assumption:
 - 1 Linearity
 - 2 Random (iid) sample
 - 3 Variation in X_i
 - 4 Zero conditional mean of the errors
 - 5 Homoskedasticity
 - 6 Errors are conditionally Normal

OLS Assumptions VI

Assumption (VI. Normality)

The population error term is independent of the explanatory variable, $u \perp\!\!\!\perp X$, and is normally distributed with mean zero and variance σ_u^2 :

$$u \sim N(0, \sigma_u^2), \text{ which implies } Y|X \sim N(\beta_0 + \beta_1 X, \sigma_u^2)$$

Note: This also implies homoskedasticity and zero conditional mean.

- Together Assumptions I–VI are the **classical linear model (CLM) assumptions**.
- The CLM assumptions imply that OLS is **BUE** (i.e. minimum variance among all linear or non-linear unbiased estimators)
- Non-normality of the errors is a serious concern in small samples. We can *partially* check this assumption by looking at the residuals (more in coming weeks)
- Variable transformations can help to come closer to normality
- Reminder: we don't need normality assumption in large samples

Sampling Distribution for $\hat{\beta}_1$

Theorem (Sampling Distribution of $\hat{\beta}_1$)

Under Assumptions I–VI,

$$\hat{\beta}_1 \sim N(\beta_1, \text{Var}[\hat{\beta}_1 | X])$$

where

$$\text{Var}[\hat{\beta}_1 | X] = \frac{\sigma_u^2}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

which implies

$$\frac{\hat{\beta}_1 - \beta_1}{\sqrt{\text{Var}[\hat{\beta}_1 | X]}} = \frac{\hat{\beta}_1 - \beta_1}{SE(\hat{\beta})} \sim N(0, 1)$$

Proof.

Given Assumptions I–VI, $\hat{\beta}_1$ is a linear combination of the i.i.d. normal random variables:

$$\hat{\beta}_1 = \beta_1 + \sum_{i=1}^n \frac{(x_i - \bar{x})}{SST_x} u_i \quad \text{where} \quad u_i \sim N(0, \sigma_u^2).$$

Any linear combination of independent normals is normal, and we can transform/standardize any normal random variable into a standard normal by subtracting off its mean and dividing by its standard deviation. □

Sampling distribution of OLS slope

- If we have Y_i given X_i is distributed $N(\beta_0 + \beta_1 X_i, \sigma_u^2)$, then we have the following at any sample size:

$$\frac{\hat{\beta}_1 - \beta_1}{SE[\hat{\beta}_1]} \sim N(0, 1)$$

- Furthermore, if we replace the true standard error with the estimated standard error, then we get the following:

$$\frac{\hat{\beta}_1 - \beta_1}{\widehat{SE}[\hat{\beta}_1]} \sim t_{n-2}$$

- The standardized coefficient follows a t distribution $n - 2$ degrees of freedom. We take off an extra degree of freedom because we had to estimate one more parameter than just the sample mean.
- All of this depends on Normal errors!

The t-Test for Single Population Parameters

- $SE[\hat{\beta}_1] = \frac{\sigma_u}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2}}$ involves the unknown population error variance σ_u^2
- Replace σ_u^2 with its unbiased estimator $\hat{\sigma}_u^2 = \frac{\sum_{i=1}^n \hat{u}_i^2}{n-2}$, and we obtain:

Theorem (Sampling Distribution of t-value)

Under Assumptions I–VI, the **t-value** for β_1 has a *t*-distribution with $n - 2$ degrees of freedom:

$$T \equiv \frac{\hat{\beta}_1 - \beta_1}{SE[\hat{\beta}_1]} \sim \tau_{n-2}$$

Proof.

The logic is perfectly analogous to the t-value for the population mean — because we are estimating the denominator, we need a distribution that has fatter tails than $N(0, 1)$ to take into account the additional uncertainty.

This time, $\hat{\sigma}_u^2$ contains two estimated parameters ($\hat{\beta}_0$ and $\hat{\beta}_1$) instead of one, hence the degrees of freedom = $n - 2$. □

Where are we?

- Under Assumptions 1-5 and in large samples, we know that

$$\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma_u^2}{\sum_{i=1}^n (X_i - \bar{X})^2}\right)$$

- Under Assumptions 1-6 and in any sample, we know that

$$\frac{\hat{\beta}_1 - \beta_1}{\widehat{SE}[\hat{\beta}_1]} \sim t_{n-2}$$

Now let's briefly return to some of the large sample properties.

Large Sample Properties: Consistency

- We just looked formally at the **small sample** properties of the OLS estimator, i.e., how $(\hat{\beta}_0, \hat{\beta}_1)$ behaves *in repeated samples* of a given n .
- Now let's take a more rigorous look at the **large sample** properties, i.e., how $(\hat{\beta}_0, \hat{\beta}_1)$ behaves *when $n \rightarrow \infty$* .

Theorem (Consistency of OLS Estimator)

Given Assumptions I-IV, the OLS estimator $\hat{\beta}_1$ is consistent for β_1 as $n \rightarrow \infty$:

$$\text{plim}_{n \rightarrow \infty} \hat{\beta}_1 = \beta_1$$

- Technical note: We can slightly relax Assumption IV:

$$E[u|X] = 0 \quad (\text{any function of } X \text{ is uncorrelated with } u)$$

to its implication:

$$\text{Cov}[u, X] = 0 \quad (X \text{ is uncorrelated with } u)$$

for consistency to hold (but not unbiasedness).

Large Sample Properties: Consistency

Proof.

Similar to the unbiasedness proof:

$$\begin{aligned}\hat{\beta}_1 &= \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2} = \beta_1 + \frac{\sum_{i=1}^n (x_i - \bar{x}) u_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \\ \text{plim } \hat{\beta}_1 &= \text{plim } \beta_1 + \text{plim } \frac{\sum_{i=1}^n (x_i - \bar{x}) u_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \quad (\text{Wooldridge C.3 Property i}) \\ &= \beta_1 + \frac{\text{plim } \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x}) u_i}{\text{plim } \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2} \quad (\text{Wooldridge C.3 Property iii}) \\ &= \beta_1 + \frac{\text{Cov}[X, u]}{\text{Var}[X]} \quad (\text{by the law of large numbers}) \\ &= \beta_1 \quad (\text{Cov}[X, u] = 0 \text{ and } \text{Var}[X] > 0)\end{aligned}$$



- OLS is inconsistent (and biased) unless $\text{Cov}[X, u] = 0$
- If $\text{Cov}[u, X] > 0$ then asymptotic bias is upward; if $\text{Cov}[u, X] < 0$ asymptotic bias is downwards

Large Sample Properties: Asymptotic Normality

- For statistical inference, we need to know the sampling distribution of $\hat{\beta}$ when $n \rightarrow \infty$.

Theorem (Asymptotic Normality of OLS Estimator)

Given *Assumptions I–V*, the OLS estimator $\hat{\beta}_1$ is asymptotically normally distributed:

$$\frac{\hat{\beta}_1 - \beta_1}{\widehat{SE}[\hat{\beta}_1]} \underset{\sim}{\text{approx.}} N(0, 1)$$

where

$$\widehat{SE}[\hat{\beta}_1] = \frac{\hat{\sigma}_u}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

with the consistent estimator for the error variance:

$$\hat{\sigma}_u^2 = \frac{1}{n} \sum_{i=1}^n \hat{u}_i^2 \xrightarrow{p} \sigma_u^2$$

Large Sample Inference

Proof.

Proof is similar to the small-sample normality proof:

$$\hat{\beta}_1 = \beta_1 + \sum_{i=1}^n \frac{(x_i - \bar{x})}{SST_x} u_i$$
$$\sqrt{n}(\hat{\beta}_1 - \beta_1) = \frac{\sqrt{n} \cdot \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x}) u_i}{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}$$

where the numerator converges in distribution to a normal random variable by CLT. Then, rearranging the terms, etc. gives you the right formula given in the theorem.

For a more formal and detailed proof, see Wooldridge Appendix 5A. □

- We need homoskedasticity (Assumption V) for this result, but **we do not need normality (Assumption VI)**.
- Result implies that **asymptotically** our usual standard errors, t-values, p-values, and CIs remain valid even without the normality assumption! We just proceed as in the small sample case where we assume normality.
- It turns out that, given Assumptions I–V, the OLS asymptotic variance is also the lowest in class (asymptotic Gauss-Markov).

Testing and Confidence Intervals

Three ways of making statistical inference out of regression:

- 1 **Point Estimation:** Consider the sampling distribution of our point estimator $\hat{\beta}_1$ to infer β_1
- 2 **Hypothesis Testing:** Consider the sampling distribution of a test statistic to test hypothesis about β_1 at the α level
- 3 **Interval Estimation:** Consider the sampling distribution of an interval estimator to construct intervals that will contain β_1 at least $100(1 - \alpha)\%$ of the time.

For 2 and 3, we need to know more than just the mean and the variance of the sampling distribution of $\hat{\beta}_1$. We need to know the full shape of the sampling distribution of our estimators $\hat{\beta}_0$ and $\hat{\beta}_1$.

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Null and alternative hypotheses review

- Null: $H_0 : \beta_1 = 0$
 - ▶ The null is the straw man we want to knock down.
 - ▶ With regression, almost always null of no relationship
- Alternative: $H_a : \beta_1 \neq 0$
 - ▶ Claim we want to test
 - ▶ Almost always “some effect”
 - ▶ Could do one-sided test, but you shouldn't
- Notice these are statements about the population parameters, not the OLS estimates.

Test statistic

- Under the null of $H_0 : \beta_1 = c$, we can use the following familiar test statistic:

$$T = \frac{\hat{\beta}_1 - c}{\widehat{SE}[\hat{\beta}_1]}$$

- As we saw in the last section, if the errors are conditionally Normal, then under the null hypothesis we have:

$$T \sim t_{n-2}$$

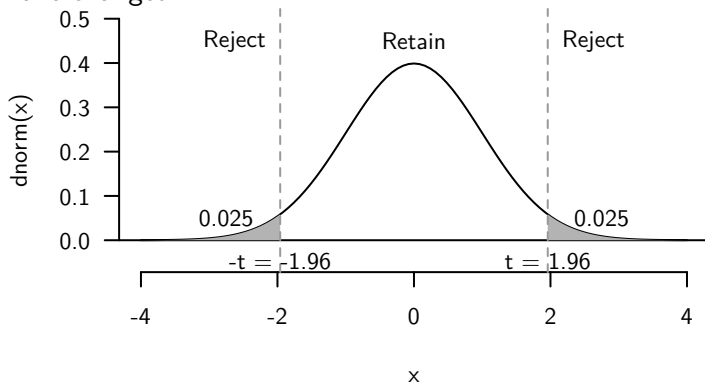
- In large samples, we know that T is approximately (standard) Normal, but we also know that t_{n-2} is approximately (standard) Normal in large samples too, so this statement works there too, even if Normality of the errors fails.
- Thus, under the null, we know the distribution of T and can use that to formulate a rejection region and calculate p-values.

Rejection region

- Choose a level of the test, α , and find rejection regions that correspond to that value under the null distribution:

$$\mathbb{P}(-t_{\alpha/2, n-2} < T < t_{\alpha/2, n-2}) = 1 - \alpha$$

- This is exactly the same as with sample means and sample differences in means, except that the degrees of freedom on the t distribution have changed.



p-value

- The interpretation of the p-value is the same: the probability of seeing a test statistic at least this extreme if the null hypothesis were true
- Mathematically:

$$\mathbb{P} \left(\left| \frac{\hat{\beta}_1 - c}{\widehat{SE}[\hat{\beta}_1]} \right| \geq |T_{obs}| \right)$$

- If the p-value is less than α we would reject the null at the α level.

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Confidence intervals

- Very similar to the approach with sample means. By the sampling distribution of the OLS estimator, we know that we can find t -values such that:

$$\mathbb{P}\left(-t_{\alpha/2, n-2} \leq \frac{\hat{\beta}_1 - \beta_1}{\widehat{SE}[\hat{\beta}_1]} \leq t_{\alpha/2, n-2}\right) = 1 - \alpha$$

- If we rearrange this as before, we can get an expression for confidence intervals:

$$\mathbb{P}\left(\hat{\beta}_1 - t_{\alpha/2, n-2}\widehat{SE}[\hat{\beta}_1] \leq \beta_1 \leq \hat{\beta}_1 + t_{\alpha/2, n-2}\widehat{SE}[\hat{\beta}_1]\right) = 1 - \alpha$$

- Thus, we can write the confidence intervals as:

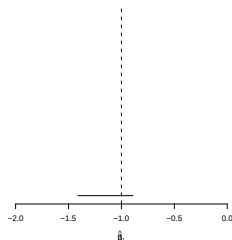
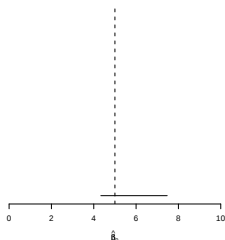
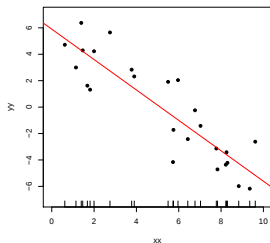
$$\hat{\beta}_1 \pm t_{\alpha/2, n-2}\widehat{SE}[\hat{\beta}_1]$$

- We can derive these for the intercept as well:

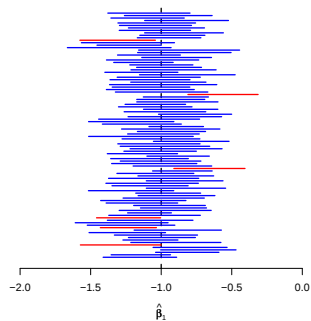
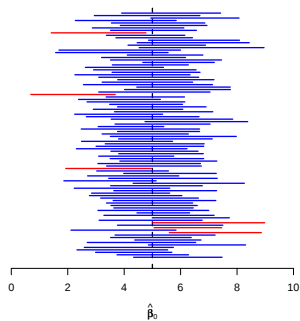
$$\hat{\beta}_0 \pm t_{\alpha/2, n-2}\widehat{SE}[\hat{\beta}_0]$$

CI's Simulation Example

Returning to our simulation example we can simulate the sampling distributions of the 95 % confidence interval estimates for $\hat{\beta}_1$ and $\hat{\beta}_0$



CI Simulation Example



Prediction error

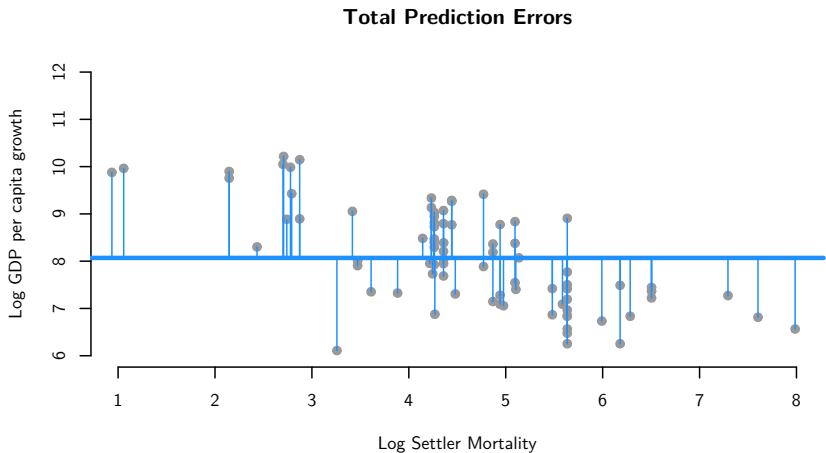
- How do we judge how well a line fits the data?
- One way is to find out how much better we do at predicting Y once we include X into the regression model.
- Prediction errors without X : best prediction is the mean, so our squared errors, or the **total sum of squares** (SS_{tot}) would be:

$$SS_{tot} = \sum_{i=1}^n (Y_i - \bar{Y})^2$$

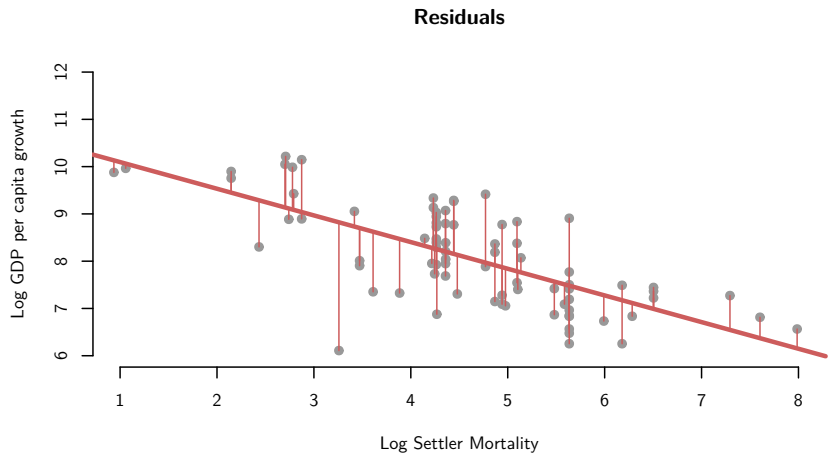
- Once we have estimated our model, we have new prediction errors, which are just the sum of the squared residuals or SS_{res} :

$$SS_{res} = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2$$

Sum of Squares



Sum of Squares



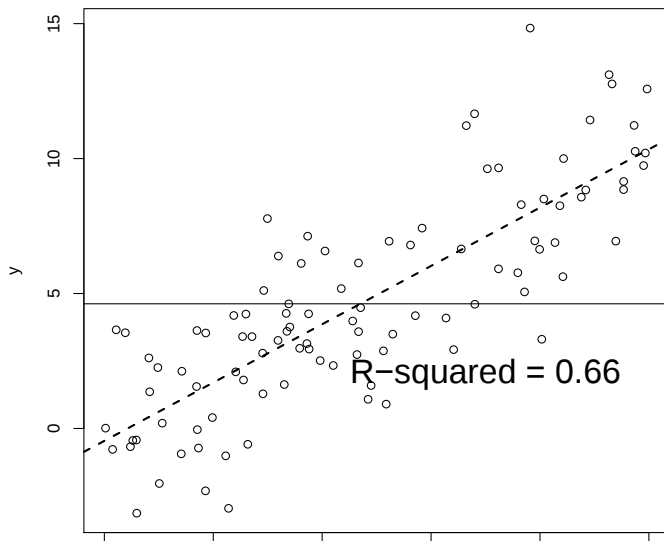
R-square

- By definition, the residuals have to be smaller than the deviations from the mean, so we might ask the following: how much lower is the SS_{res} compared to the SS_{tot} ?
- We quantify this question with the **coefficient of determination** or R^2 . This is the following:

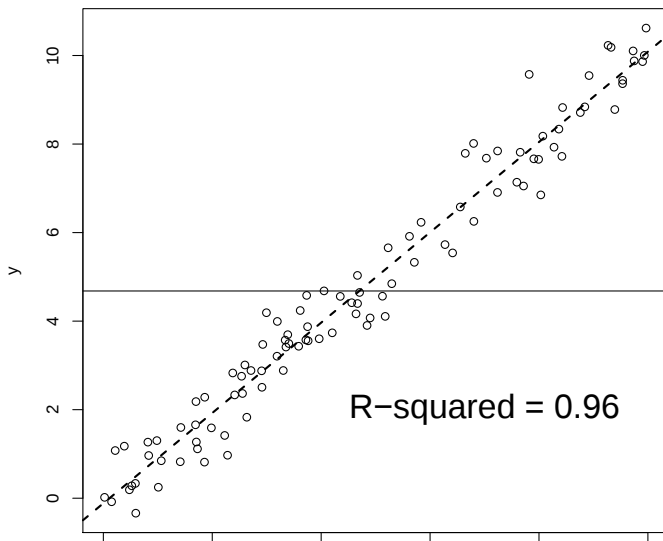
$$R^2 = \frac{SS_{tot} - SS_{res}}{SS_{tot}} = 1 - \frac{SS_{res}}{SS_{tot}}$$

- This is the fraction of the total prediction error eliminated by providing information on X .
- Alternatively, this is the fraction of the variation in Y is “explained by” X .
- $R^2 = 0$ means no relationship
- $R^2 = 1$ implies perfect linear fit

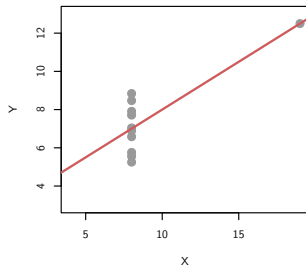
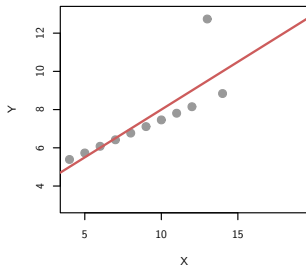
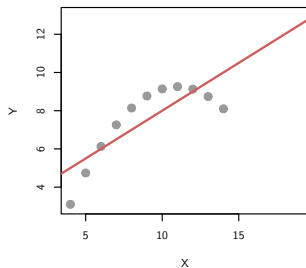
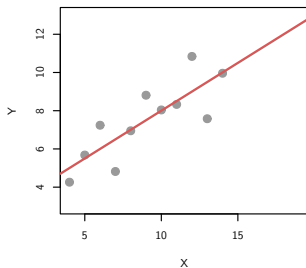
Is R-squared useful?



Is R-squared useful?

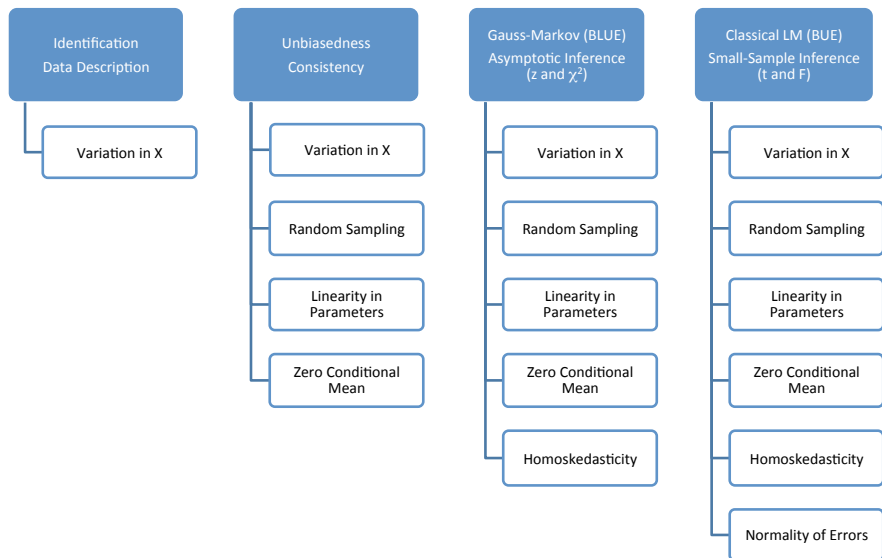


Is R-squared useful?



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OLS Assumptions Summary



What Do the Regression Coefficients Mean Substantively?

- So far, we have learned the **statistical properties** of the OLS estimator
- However, these properties do not tell us what **types of inference** we can draw from the estimates

Three types of inference:

① Descriptive inference:

- ▶ Summarizing sample data by drawing the “best fitting” line
- ▶ No inference about the underlying population intended
- ▶ Assumption required: III (variation in X) only

② Predictive inference:

- ▶ Inference about a **new observation** coming from the same population
- ▶ Example: Wage (Y) and education (X):
“What’s my best guess about the wage of a new worker who only has high school education?”
- ▶ Assumptions required: III and II (random sampling)
- ▶ Assumptions desired: I (linearity)

What Do the Regression Coefficients Mean Substantively?

③ Causal inference:

- ▶ Inference about **counterfactuals**, i.e. hypothetical **interventions** to the same units
- ▶ Example: Wage (Y) and education (X):
“What would the wages of a non-college educated worker be if we sent them to college?”
- ▶ We will come back to this in the last few weeks.

OLS as a Best Linear Predictor (Review of BLUE)

- Suppose that we want to predict the values of Y given observed X values
- Suppose further that we've decided to *use* a linear predictor $\hat{\beta}_0 + \hat{\beta}_1 X$ (but not necessarily *assume* a true linear relationship in the population)
- How to choose a good predictor? A popular criterion is **mean squared error**:

$$MSE = E[(Y_i - \hat{Y}_i)^2] = E[(Y_i - \hat{\beta}_0 - \hat{\beta}_1 X_i)^2] = E[\hat{u}_i^2]$$

The smaller a predictor makes MSE , the better.

- Now, note that the sample version of $MSE = \frac{1}{n} \sum_{i=1}^n \hat{u}_i^2$
- Recall how we got the OLS estimator; we *minimized* $\sum_{i=1}^n \hat{u}_i^2$!
- This implies that OLS is the **best linear predictor** in terms of MSE
- Which assumptions did we use to get this result?
 - ▶ Needed: Assumptions II (random sampling) and III (variation in X)
 - ▶ *Not* needed: Assumptions I (linearity) and IV (zero cond. mean)
- Note that Assumption I would make OLS the **best**, not just best linear, **predictor**, so it is certainly desired

State Legislators and African American Population

Interpretations of increasing quality:

```
> summary(lm(beo ~ bpop, data = D))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-1.31489	0.32775	-4.012	0.000264 ***
bpop	0.35848	0.02519	14.232	< 2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.317 on 39 degrees of freedom

Multiple R-squared: 0.8385, Adjusted R-squared: 0.8344

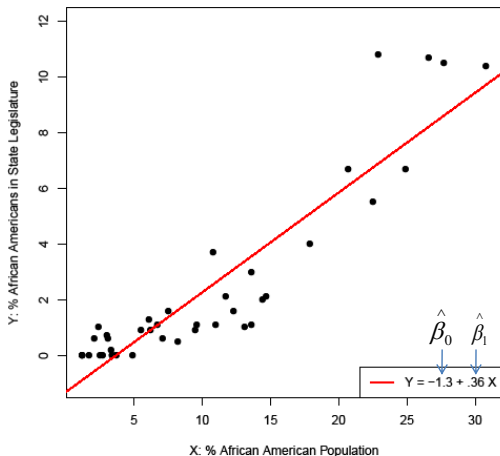
F-statistic: 202.6 on 1 and 39 DF, p-value: < 2.2e-16

“In states where an additional .01 proportion of the population is African American, we observe on average .035 proportion more African American state legislators (between .03 and .04 with 95% confidence).”

(still not perfect, the best will be subject matter specific. is fairly clear it is non-causal, gives uncertainty.)

Graphical

Graphical presentations are often the most informative. We will talk more about them later in the semester.



Ground Rules: Interpretation of the Slope

I almost didn't include the last example in the slides. It is **hard** to give ground rules that cover all cases. Regressions are a part of marshaling evidence in an argument which makes them naturally specific to context.

- ① Give a short, but precise interpretation of the association using interpretable **language** and **units**
- ② If the association has a **causal** interpretation explain why, otherwise do not imply a causal interpretation.
- ③ Provide a meaningful sense of **uncertainty**
- ④ Indicate the **practical** significance of the finding for your argument.

Next Week

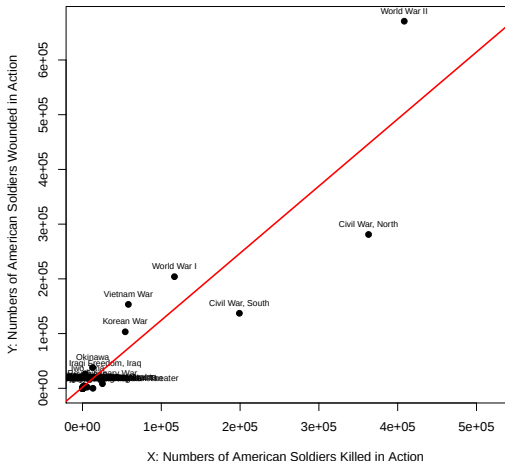
- OLS with two regressors
- Omitted Variables and Multicollinearity
- Dummy variables, interactions, polynomials
- Reading:
 - ▶ Optional Fox Chapters 5-7
 - ▶ For more on logs, Gelman and Hill (2007) pg 59-61 is nice

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Fun with Non-Linearities

- The linear regression model *can* accommodate non-linearity in X (but not in β)
- We do this by first **transforming** X appropriately
- A useful transformation when variables are positive and right-skewed is the (natural) logarithm
- The log transformation changes the interpretation of β_1 :
 - ▶ Regress $\log(Y)$ on $X \rightarrow \beta_1$ approximates **percent increase** in Y associated with one unit increase in X
 - ▶ Regress Y on $\log(X) \rightarrow \beta_1$ approximates increase in Y associated with a **percent increase** in X
 - ▶ Note that these approximations work only for small increments
 - ▶ In particular, they do not work when X is a discrete random variable

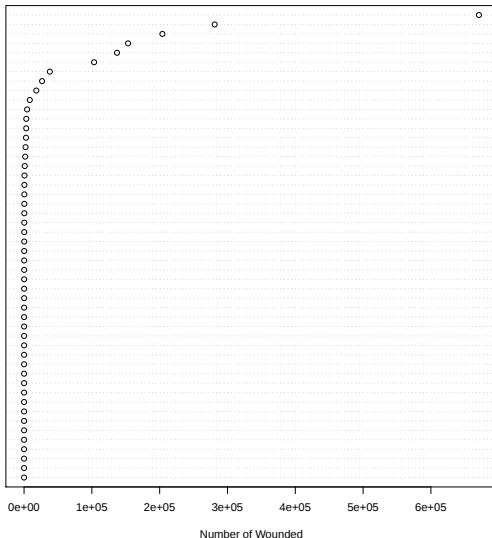
Example from the American War Library



$\hat{\beta}_1 = 1.23 \rightarrow$ One additional soldier killed predicts 1.23 additional soldiers wounded on average

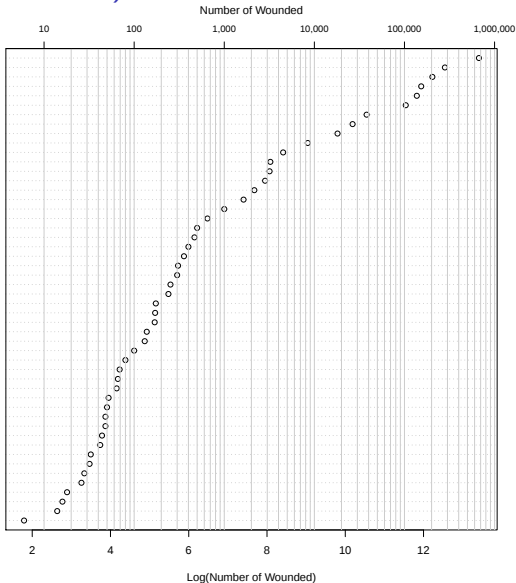
Wounded (Scale in Levels)

World War II
 Civil War, North
 Vietnam War I
 Civil War, South
 Korean War
 Okinawa
 Operation Iraqi Freedom, Iraq
 Iwo Jima
 Revolutionary War
 War of 1812
 Aleutian Campaign
 D-Day
 Philippines War
 Indian Wars
 Spanish American War
 Terrorism, World Trade Center
 Yemen, USS Cole
 Terrorism Khobar Towers, Saudi Arabia
 Persian Gulf
 Terrorism Oklahoma City
 Persian Gulf, Op Desert Shield/Storm
 Russia North Expedition
 Moro Campaigns
 China Boxer Rebellion
 Panama
 Dominican Republic
 Israel Attack/USS Liberty
 Lebanon
 Texas War Of Independence
 South Korea
 Grenada
 China Yangtze Service
 Mexico
 Nicaragua
 Barbary Wars
 Russia Siberia Expedition
 Dominican Republic
 China Civil War
 Terrorism Riyadh, Saudi Arabia
 North Atlantic Naval War
 Franco-Amer Naval War
 Operation Enduring Freedom, Afghanistan
 Mexican War
 Operation Enduring Freedom, Afghanistan Theater
 Haiti
 Texas Border Cortina War
 Nicaragua
 Italy Trieste
 Japan

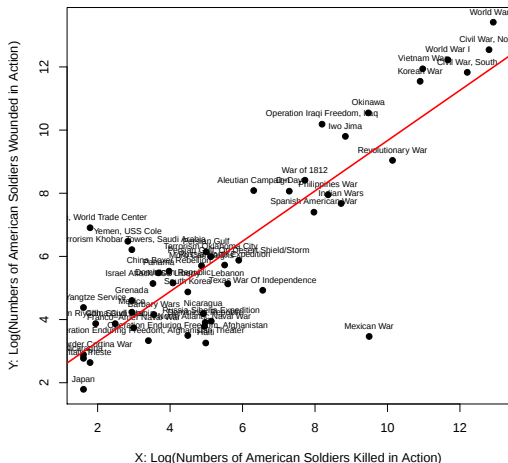


Wounded (Logarithmic Scale)

World War II
 Civil War, North
 World War I
 Vietnam War
 Civil War, South
 Korean War
 Okinawa
 Operation Iraqi Freedom, Iraq
 Iwo Jima
 Revolutionary War
 War of 1812
 Aleutian Campaign
 D-Day
 Philippines War
 Indian Wars
 Spanish American War
 Terrorism, World Trade Center
 Yemen, USS Cole
 Terrorism Khobar Towers, Saudi Arabia
 Persian Gulf
 Terrorism Oklahoma City
 Persian Gulf, Op Desert Shield/Storm
 Russia North Expedition
 Moro Campaigns
 China Boxer Rebellion
 Panama
 Dominican Republic
 Israel Attack/USS Liberty
 Lebanon
 Texas War Of Independence
 South Korea
 Grenada
 China Yangtze Service
 Mexico
 Nicaragua
 Barbary Wars
 Russia Siberia Expedition
 Dominican Republic
 China Civil War
 Terrorism Riyadh, Saudi Arabia
 North Atlantic Naval War
 Franco-Amer Naval War
 Operation Enduring Freedom, Afghanistan
 Mexican War
 Operation Enduring Freedom, Afghanistan Theater
 Haiti
 Texas Border Cortina War
 Nicaragua
 Italy Trieste
 Japan



Regression: Log-Log



$\hat{\beta}_1 = 0.797 \rightarrow$ A percent increase in deaths predicts 0.797 percent increase in the wounded on average

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Why r^2 ?

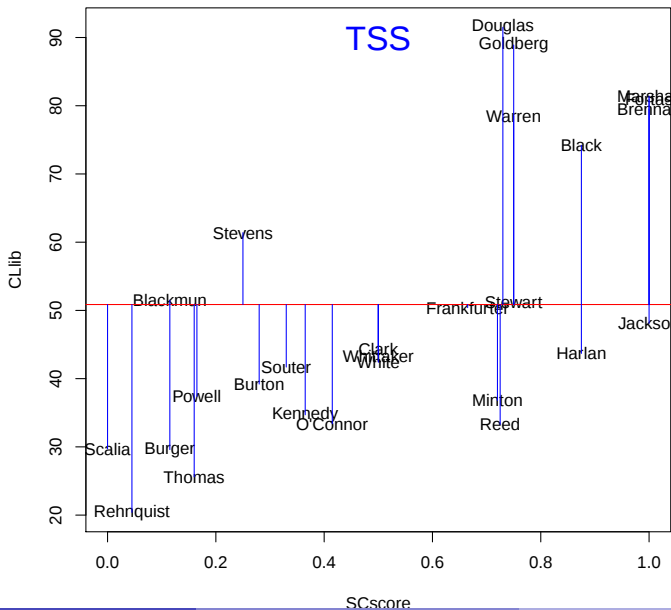
To calculate r^2 , we need to think about the following two quantities:

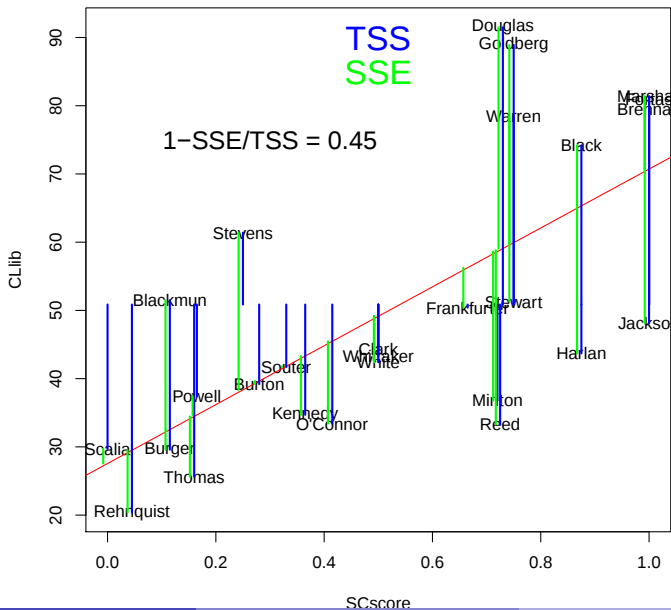
- ① TSS: Total sum of squares
- ② SSE: Sum of squared errors

$$TSS = \sum_{i=1}^n (y_i - \bar{y})^2.$$

$$SSE = \sum_{i=1}^n u_i^2.$$

$$r^2 = 1 - \frac{SSE}{TSS}.$$





Derivation

$$\begin{aligned}\sum_{i=1}^n (y_i - \bar{y})^2 &= \sum_{i=1}^n \{\hat{u}_i + (\hat{y}_i - \bar{y})\}^2 \\&= \sum_{i=1}^n \{\hat{u}_i^2 + 2\hat{u}_i(\hat{y}_i - \bar{y}) + (\hat{y}_i - \bar{y})^2\} \\&= \sum_{i=1}^n \hat{u}_i^2 + 2 \sum_{i=1}^n \hat{u}_i(\hat{y}_i - \bar{y}) + \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 \\&= \sum_{i=1}^n \hat{u}_i^2 + \sum_{i=1}^n (\hat{y}_i - \bar{y})^2 \\TSS &= SSE + RegSS\end{aligned}$$

Coefficient of Determination

We can divide each side by the TSS:

$$\frac{SSE}{TSS} + \frac{RegSS}{TSS} = \frac{TSS}{TSS}$$

$$\frac{SSE}{TSS} + \frac{RegSS}{TSS} = 1$$

$$\frac{RegSS}{TSS} = 1 - \frac{SSE}{TSS} = r^2$$

r^2 is a measure of how much of the variation in Y is accounted for by X .

References

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Wooldridge, Jeffrey. 2000. *Introductory Econometrics*. New York: South-Western.